
IDENTIFYING BROAD AND NARROW FINANCIAL RISK FACTORS VIA CONVEX OPTIMIZATION: PART II

Alexander D. Shkolnik

ads2@berkeley.edu

MMDS Workshop. June 22, 2016.

joint with Jeffrey Bohn and Lisa Goldberg.

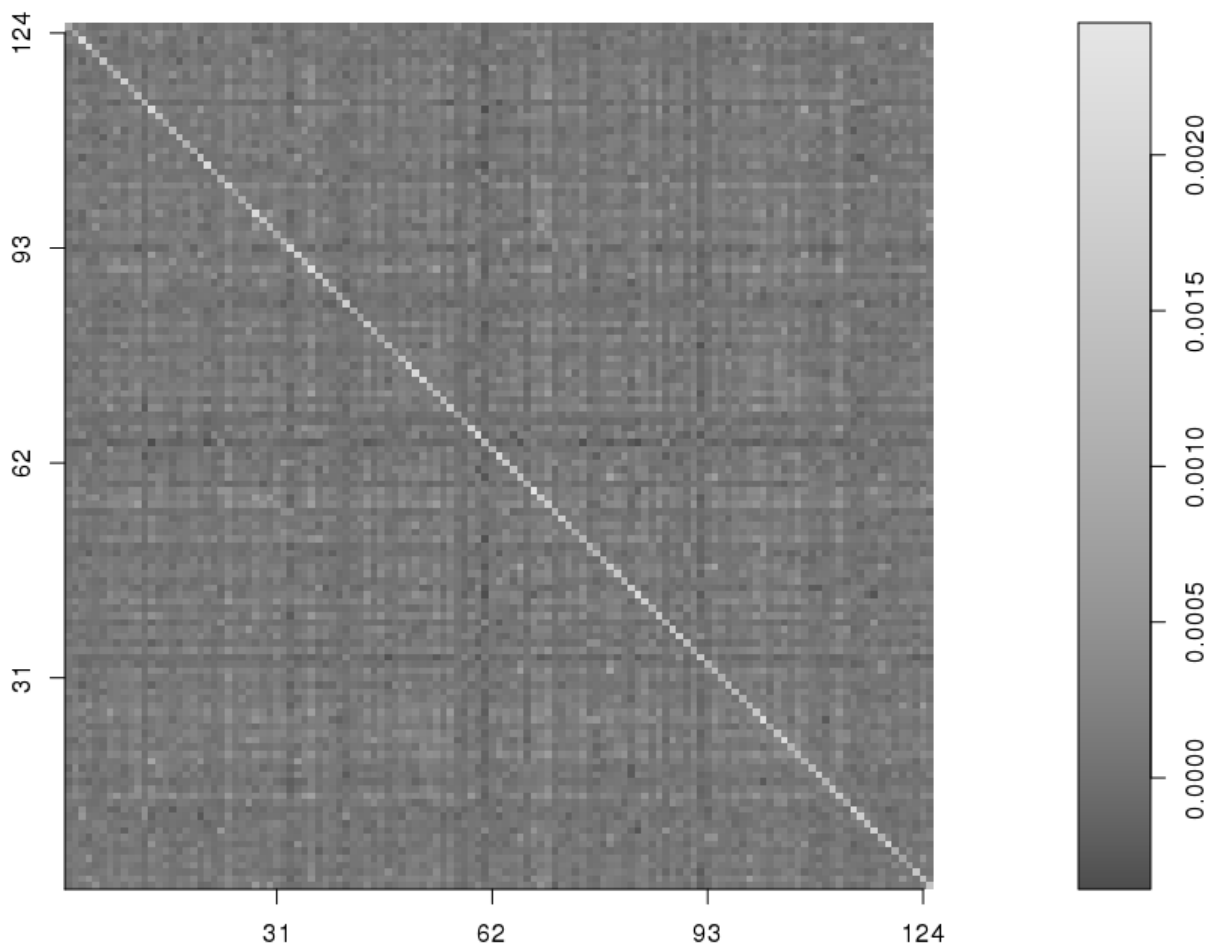
Overview

- Simulations (synthetic data).
 - Performance metrics (finance).
 - benchmark: **PCA**.
 - Theoretical considerations for $N > T$.
-

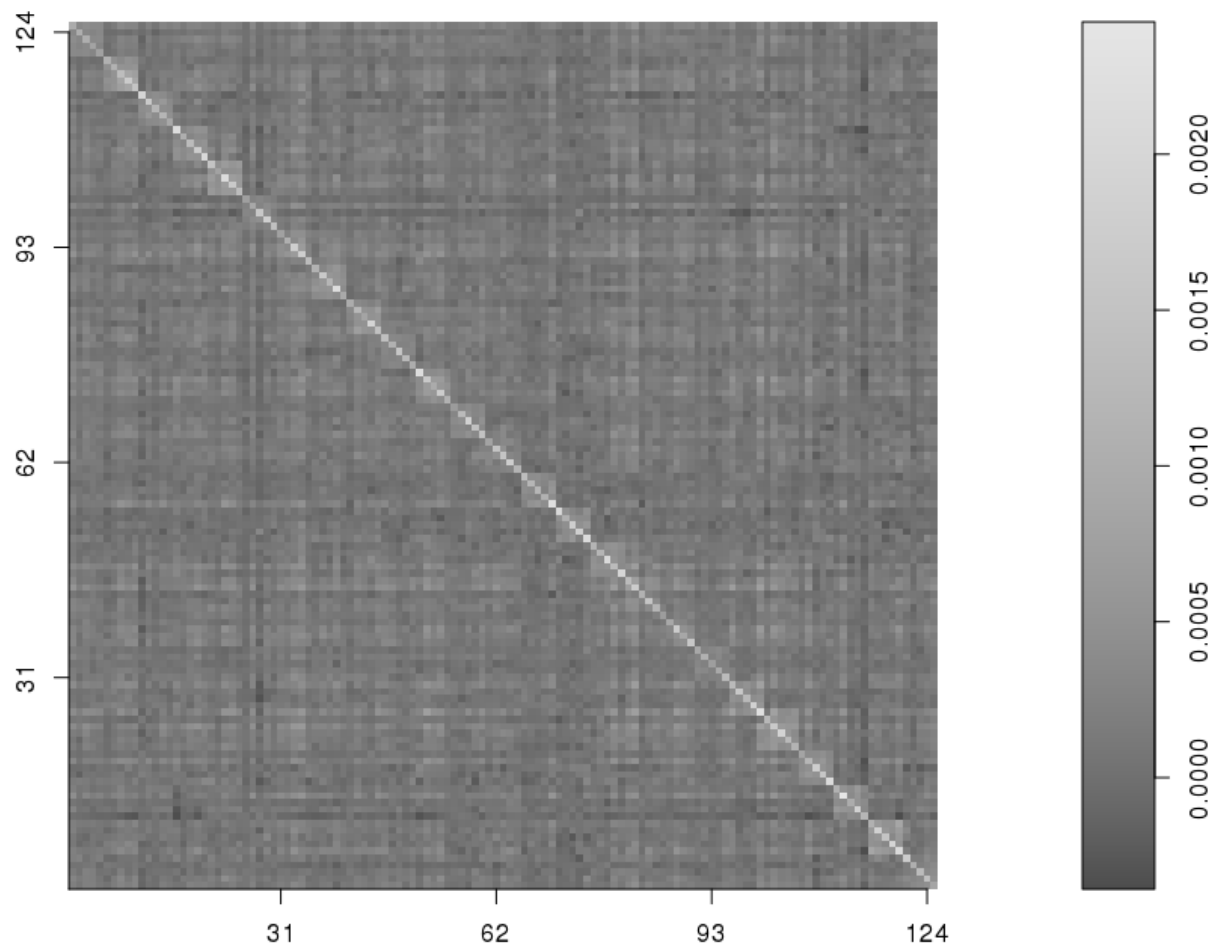
Simulations (synthetic data)

- $N = 125$ and $N = 500$ securities.
 - $T = 250$ observations (one year of daily data).
 - $K = 2$ broad factors (16% ann. vol., 4% ann. vol.)
 - $\kappa \approx N / \log N$ and $\kappa \approx \sqrt{N}$ countries.
 - 400 simulations (realizations of a sample covariance matrix).
 - **Convex program parameters:** $\theta = (\gamma, \lambda)$,
 - $\gamma = 1 / \sqrt{N}$ (see Candès, Li, Ma & Wright (2011)),
 - λ set to control the number of recovered broad factors.
-

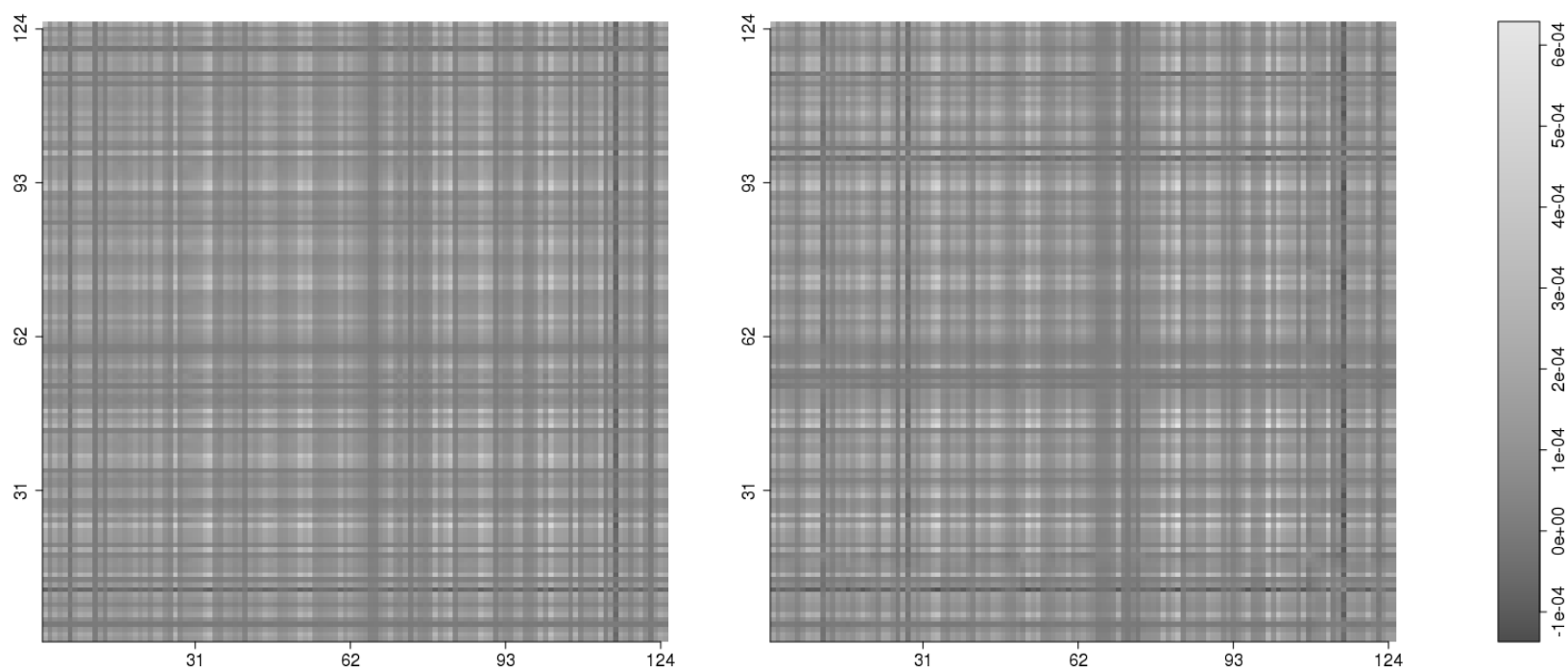
Input to algorithm ($N = 125, T = 250$)



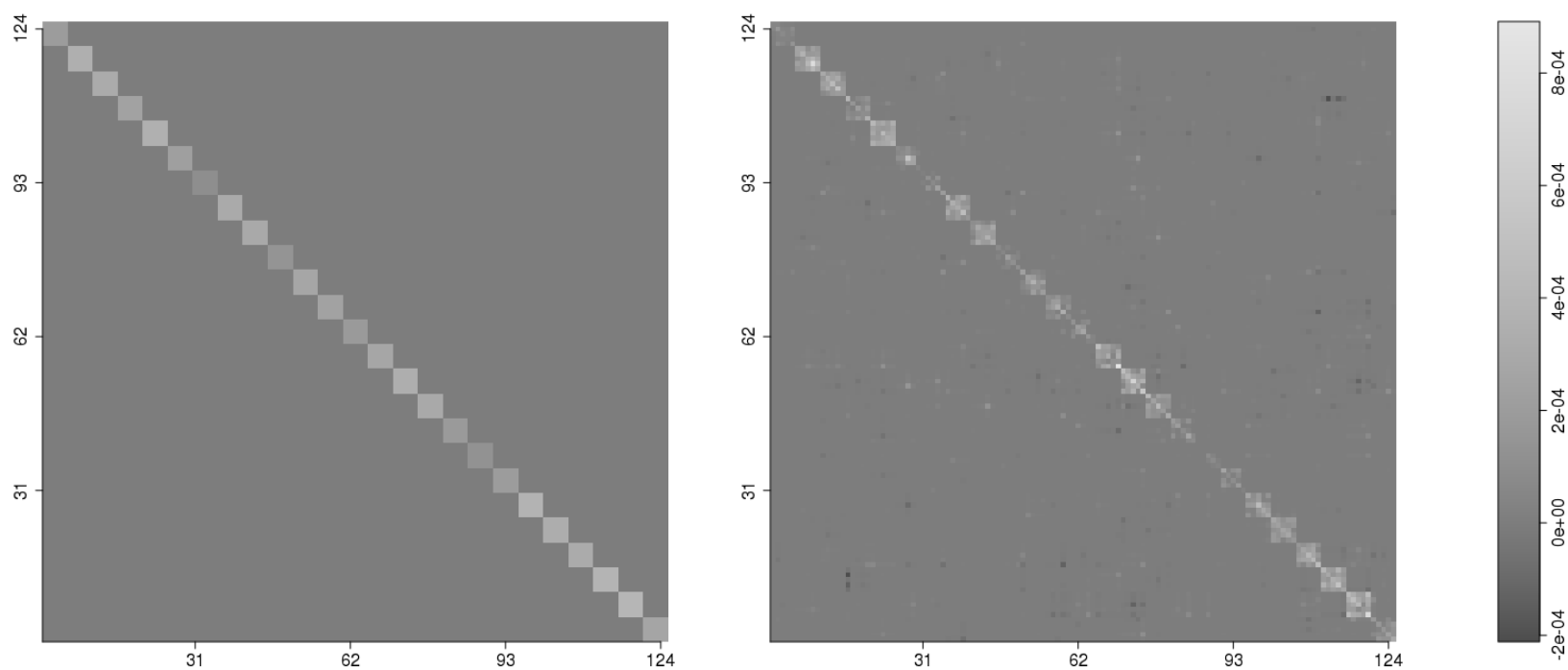
Ordered by country



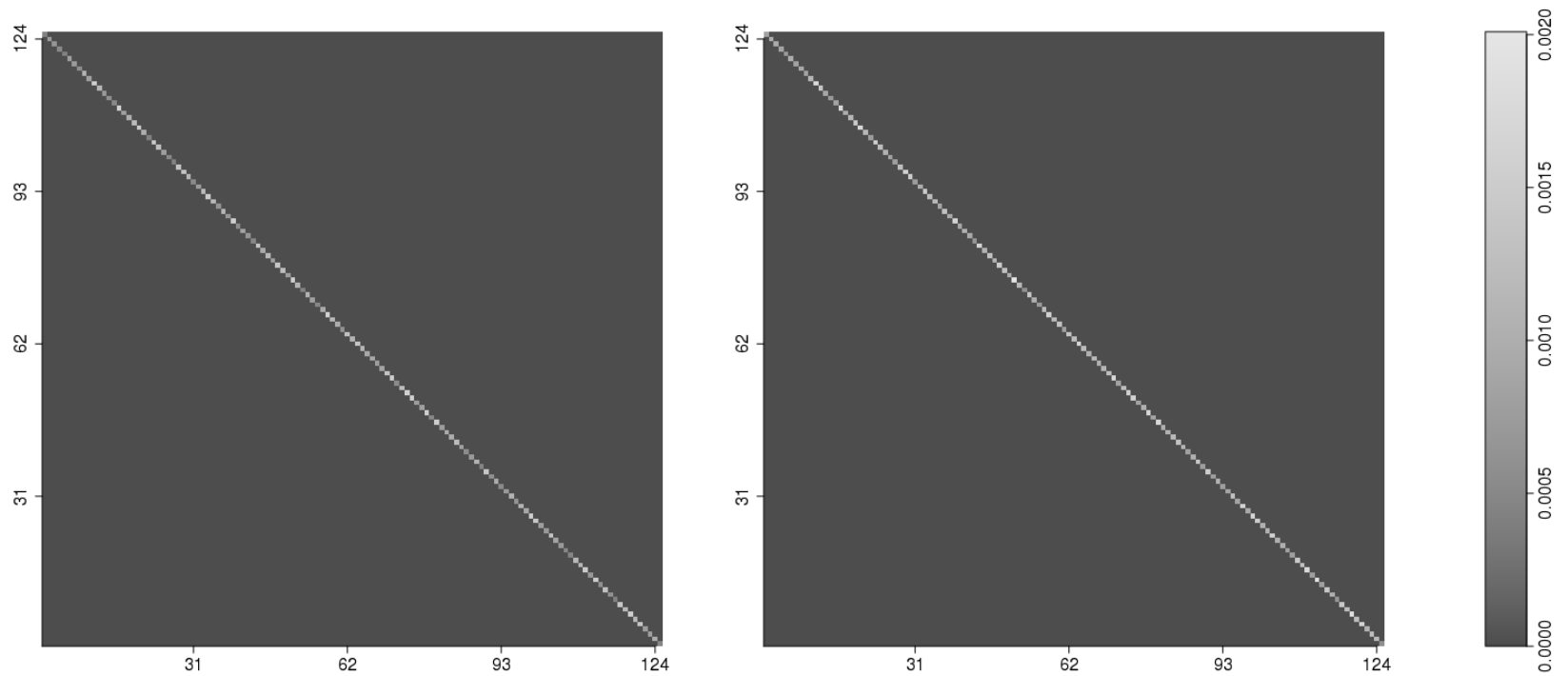
True vs recovered broad factor matrix L_K



True vs recovered narrow factor matrix L_K



True vs recovered specific risk matrix



Measuring performance on simulated data

- Typically, for estimator $\Sigma(\theta)$, consider $\|\Sigma(\theta) - \Sigma\|$ for some norm.
- Instead, consider the variance of returns to portfolio w .

$$\mathbf{Var}_{\Sigma}(w) = w^{\top} \Sigma w \quad (1)$$

- The portfolio **risk forecasting ratio** is computed as

$$\mathcal{R}_{\Sigma}(w | \theta) = \frac{\mathbf{Var}_{\Sigma(\theta)}(w)}{\mathbf{Var}_{\Sigma}(w)}. \quad (2)$$

- Ratios \mathcal{R}_{L_K} , $\mathcal{R}_{L_{\kappa}}$ and \mathcal{R}_{Δ} are defined analogously.
 - Other estimators: $K(\theta)$ and $\kappa(\theta)$.
-

Test portfolios

- Let $e = (1, \dots, 1)^\top$. **Equally weighted** portfolio is given by

$$w = e/N \quad (3)$$

- **Minimum variance** (long only) portfolio is the solution $w(\theta)$ of

$$\min_w w^\top \Sigma(\theta) w \quad (4)$$

$$\text{subject to } w^\top e = 1, \quad (5)$$

$$w \geq 0. \quad (6)$$

- **Minimum variance** (long/short): solves (4) subject to $w^\top e = 1$,

$$w(\theta) = \frac{\Sigma^{-1}(\theta) e}{e^\top \Sigma^{-1}(\theta) e}. \quad (7)$$

Number of factors: equally weighted portfolio

N	Metric	DSLRLR: ($\lambda \times 10^3$)		PCA	Ground Truth K, κ
		6.4	12.8		
500	$K(\theta)$	8.51	2.45	2.00	2
	(stdev)	(0.83)	(0.44)	(0.00)	
	$\kappa(\theta)$	20.3	22.4	n/a	22
	(stdev)	(0.21)	(0.86)	(0.00)	

★ $T = 250$ (observations).

Performance: equally weighted portfolio

N	Metric	DSLRLR: ($\lambda \times 10^3$)		PCA	Ground Truth
		6.4	12.8		Ann. Vol.
500	\mathcal{R}_Σ	1.00	0.96	1.06	0.17
	(stdev)	(0.08)	(0.08)	(0.08)	
	\mathcal{R}_{L_K}	0.97	0.94	1.18	0.16
	(stdev)	(0.09)	(0.09)	(0.10)	
	\mathcal{R}_{L_κ}	1.05	0.68	n/a	0.05
	(stdev)	(0.09)	(0.05)	(0.00)	
	\mathcal{R}_Δ	0.89	1.12	1.16	0.02
	(stdev)	(0.01)	(0.01)	(0.01)	

Performance: optimized portfolio

N	Metric	DSLRLR: ($\lambda \times 10^3$)		PCA	Ground Truth
		6.4	12.8		Ann. Vol.
500	\mathcal{R}_Σ	0.97	0.88	0.81	0.11
	(stdev)	(0.06)	(0.06)	(0.07)	
	\mathcal{R}_{L_K}	0.86	0.70	0.83	0.07
	(stdev)	(0.08)	(0.10)	(0.13)	
	\mathcal{R}_{L_κ}	0.95	0.48	n/a	0.04
	(stdev)	(0.10)	(0.07)	(0.00)	
\mathcal{R}_Δ	0.96	1.22	1.31	0.08	
(stdev)	(0.02)	(0.03)	(0.05)		

Summary of findings

- **SLRD** outperforms classical **PCA** on risk forecasts.
 - Accurate sparse component recovery implies:
 - accurate \mathcal{R}_Δ forecast,
 - accurate $\kappa(\theta)$ estimates.
 - Inaccurate $K(\theta)$ estimates **does not** imply poor risk forecast.
 - $N > T$?
-

$\kappa(\theta)$ estimates ($N > T$)

- With high probability **CPW^a** recovers correctly the
 - rank of \mathcal{L} and sparsity support of \mathcal{S} .
 - **Rate of convergence** $\sqrt{N/T}$ arises from bounds on

$$\|\mathcal{L} - \mathcal{L}(\theta)\|_2 \text{ and } \|\mathcal{S} - \mathcal{S}(\theta)\|_{\ell_\infty}. \quad (8)$$

- The error on $\|\mathcal{S} - \mathcal{S}(\theta)\|_{\ell_\infty}$ may be improved to order

$$\sqrt{\frac{\log N}{T}}. \quad (9)$$

- When sparsity pattern is correct, $\kappa(\theta)$ is correct.

^a(Chandrasekaran, Parrilo & Willsky 2012)

$K(\theta)$ estimates ($N > T$)

- With high probability **CPW^a** recovers correctly the
 - rank of \mathcal{L} and sparsity support of S .
 - **Rate of convergence** $\sqrt{N/T}$ arises from bounds on

$$\|\mathcal{L} - \mathcal{L}(\theta)\|_2 \quad (10)$$

- **Cannot** estimate number of broad factors accurately for $N > T$.
- Numerical experiments suggest $\|\mathcal{L} - \mathcal{L}(\theta)\|_{\ell_\infty}$ may be order

$$\sqrt{\frac{\log N}{T}}. \quad (11)$$

^a(Chandrasekaran et al. 2012)

Risk forecasts

- **Equally weighted:** accuracy of portfolio risk forecast depends on

$$\left| \mathcal{R}_\Sigma(w | \theta) - 1 \right| \lesssim \|\mathcal{L}(\theta) - \mathcal{L}\|_{\ell_\infty} \wedge \|\mathcal{S}(\theta) - \mathcal{S}\|_{\ell_\infty} \quad (12)$$

$$= O\left(\sqrt{\frac{\log N}{T}}\right). \quad (13)$$

- **Optimized:** same bound holds with different constants.
 - Similar for factor and specific risk forecasts.
 - Support for accuracy of risk forecasts in $\log(N)/T \downarrow 0$ regime.
-

Conclusions

- The approach shows promise in extracting narrow factors that highlight country/industry relationships in data.
 - Financial applications require a reformulation of available low rank and sparse decomposition methods.
 - **low-rank + sparse + diagonal** decomposition,
 - sparse eigenvectors (narrow factors).
 - **Finance oriented performance metrics** may lead to an alternative analysis of estimator consistency.
-

Ongoing & future work

- Theoretical performance guarantees for risk forecasting ratios.
 - Data-driven methods of selecting optimal parameters $\theta = (\lambda, \gamma)$.
 - Alternative convex programs guided by performance metrics.
 - Scale algorithms to tens of thousands of securities.
-

Questions.
(ads2@berkeley.edu)

\mathcal{R}_Δ (specific risk) forecast

- **MTFA^a** decomposes a given matrix S into a low rank $L_\kappa = XGX^\top$ component and a diagonal component Δ exactly if

$$\max_{x \in \mathcal{X} / \{0_N\}} \frac{\|x\|_\infty}{\|x\|_2} < 1/2. \quad (14)$$

where \mathcal{X} is the space spanned by the narrow factor exposures in X .

- A sufficient condition for (14) is

$$x_i^2 < \sum_{i \neq j} x_j^2 \quad \text{for all } i \text{ and } x \in \mathcal{X}. \quad (15)$$

- **Robust recovery for factors that are not too narrow.**

^a(Saunderson, Chandrasekaran, Parrilo & Willsky 2012)

References

- Candès, Emmanuel J, Xiaodong Li, Yi Ma & John Wright (2011), ‘Robust principal component analysis?’, *Journal of the ACM (JACM)* **58**(3), 11.
- Chandrasekaran, Venkat, Pablo A Parrilo & Alan S Willsky (2012), ‘Latent variable graphical model selection via convex optimization’, *The Annals of Statistics (with discussion)* **40**(4), 1935–1967.
- Saunderson, James, Venkat Chandrasekaran, Pablo A Parrilo & Alan S Willsky (2012), ‘Diagonal and low-rank matrix decompositions, correlation matrices, and ellipsoid fitting’, *SIAM Journal on Matrix Analysis and Applications* **33**(4), 1395–1416.
-